

# PROJECTION INFERENCE FOR SET-IDENTIFIED SVARS

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## INTRODUCTION: SET-ID. SVARs

- ★ SVAR: Theoretical restrictions  $\mathcal{R}$  imposed on a VAR.  
(Sims [1980, 1986])

$$Y_t = A_1 Y_{t-1} + \dots + A_p Y_{t-p} + \eta_t, \quad \Sigma = \mathbb{E}[\eta_t \eta_t']$$

- ★ Goal of the restrictions:  $(A_1, \dots, A_p, \Sigma) \mapsto_{\mathcal{R}} \text{IRF}_{k,i,j}$   
(response of variable  $i$  to a  $j$ -th 'structural shock' at horizon  $k$ )
- ★ Map ' $\mapsto_{\mathcal{R}}$ ' can be 1-to-1 (point id.) or 1-to-many (set id.).  
(set i.d. SVARs have become popular in applied macro work)
- ★ Common practice: set-identify SVARs with  $\geq$  / = restrictions.  
(Faust [1998]; Canova and De Nicolò [2002]; Uhlig[2005])

## MOTIVATION

- ★ Most empirical studies report Bayesian *credible* sets for  $\text{IRF}_{k,i,j}$  (Bayesian Inference depends on the specification of prior beliefs)
- ★ Practical concern: prior beliefs are not ‘dominated’ by the data (results are sensitive to the choice of priors even if  $T \rightarrow \infty$ )
- ★ Theoretical Critique: Coverage and ‘Robust’ credibility  $\rightarrow 0$ . (as  $T \rightarrow \infty$ ; Moon & Schorfheide [2012], Kitagawa [2012])
- ★ Recent work on non-Bayesian Inference for set-id. SVARs. (MSG [2013]-Freq. Inference; GK [2014]-Robust Bayes)
- ★ Is there a simple way to conduct inference in set-id. SVARs? (that pleases both a frequentist and a robust Bayesian, and that is general and computationally feasible?)

## DESCRIPTION OF THE INFERENCE PROBLEM

$$\text{IRF}_{k,i,j} \in \mathcal{I}_{k,i,j}^{\mathcal{R}}(\mu) \subseteq [\underline{v}_{k,i,j}(\mu), \bar{v}_{k,i,j}(\mu)],$$
$$\mu \equiv (A, \Sigma).$$

## THIS PAPER

- ★ Studies the properties of ‘projection inference’ for set i.d. SVARs. (Scheffé [1953]; Dufour [1990]; Dufour and Taamouti [2005])
- ★ We collect  $\text{IRF}_{k,i,j}$ ’s in a  $1 - \alpha$  Wald Ellipsoid for  $\mu \equiv (A, \Sigma)$ . (that is, we ‘project’ a nominal  $1 - \alpha$  Wald Ellipsoid)
- ★ Strategy: focus on the endpoints of the identified set for  $\text{IRF}_{k,i,j}$ . (the maximum and minimum response,  $\bar{v}_{k,i,j}(\mu)$ ,  $\underline{v}_{k,i,j}(\mu)$ )

$$\left[ \inf_{\mu \in \text{CS}_T(1-\alpha)} \underline{v}_{k,i,j}(\mu), \sup_{\mu \in \text{CS}_T(1-\alpha)} \bar{v}_{k,i,j}(\mu) \right]$$

- ★ Our projection region has coverage and RB credibility  $\geq 1 - \alpha$ . (for any vector of IRFs! Thus providing simultaneous inference)

## PROS & CONS

### Pros:

- ★ **Generality:** can handle the typical application in applied work (+/0 restrictions on IRFs, long-run restrictions, elasticity bounds)
- ★ **Feasibility:** solve two nonlinear optimization problems per IRF $_{k,i,j}$  (we use state-of-the-art solution algorithms for these problems)

### Cons:

- ★ **Projection is conservative** for a frequentist and a Robust Bayesian (coverage and robust credibility are strictly above  $1-\alpha$ .)
- ★ **We 'calibrate' projection** to remove the excess of Robust Cred. (=  $1 - \alpha$  and not  $> 1 - \alpha$ . Calibration based on KMS[2016])

## OUTLINE

1. MODEL AND MAIN DEFINITIONS
2. ASSUMPTIONS AND RESULTS
3. IMPLEMENTATION AND ILLUSTRATIVE EXAMPLE
4. CONCLUSION

# 1. Model and Main Definitions



## SVAR(P)

- ★ Structural VAR for the  $n$ -dimensional vector  $Y_t$ :

$$Y_t = A_1 Y_{t-1} + \dots + A_p Y_{t-p} + B \varepsilon_t, \quad \Sigma \equiv BB'$$

- ★ Vector of reduced form parameters is:

$$\mu = (\text{vec}(A_1, A_2, \dots, A_p)', \text{vech}(\Sigma)')' \in \mathbb{R}^d$$

- ★ Coefficients of the Structural Impulse Response Function:

$$\text{IRF}^H = \{\text{IRF}_{k_h, i_h, j_h}(A, B)\}_{h=1}^H, \quad \text{IRF}_{k_h, i_h, j_h}(A, B) = \underbrace{e'_{i_h} C_{k_h}(A)}_{1 \times n} B_{j_h}.$$

- ★ Interested in simultaneous inference about  $\lambda^H \equiv \text{IRF}^H$ .  
(Inoue and Kilian [2013,2016] and Lütkepohl et. al [2016])

## RESTRICTIONS $\mathcal{R}(\mu)$ ON $B$

- ★ Identified set for  $\lambda^H$ :

$$\mathcal{I}_H^{\mathcal{R}}(\mu) \equiv \left\{ \lambda \in \mathbb{R}^H \mid \lambda_h = \text{IRF}_{k_h, i_h, j_h} \text{ s.t. } BB' = \Sigma, B \in \mathcal{R}(\mu), \forall h \right\}$$

- ★  $\pm/0$  restrictions on IRFs:  $e_{j'}' C_{k'}(A) B_{j'} \geq 0$   
(e.g. Sims [1980], Uhlig [2005])
- ★  $\pm/0$  long-run restrictions:  $e_{j'}' (\mathbb{I}_n - A(1))^{-1} B_{j'} \geq 0$   
(e.g. Blanchard, Quah [1989], Galí [1999])
- ★ Elasticity bounds:  $(e_{j'}' B_{j'}) / (e_{i'}' B_{j'}) \in [c, d]$   
(e.g. Kilian, Murphy [2012], Baumeister, Hamilton [2015])

## BOUNDS ON THE IDENTIFIED SET: MAX AND MIN RESPONSE

- ★ The endpoints of the identified set for each IRF $_{k,i,j}$ :

$$\bar{v}_{k,i,j}(\mu) \equiv \sup_B \text{IRF}_{k,i,j}(A, B) \quad \text{s.t. } BB' = \Sigma, B \in \mathcal{R}(\mu)$$

$$\underline{v}_{k,i,j}(\mu) \equiv \inf_B \text{IRF}_{k,i,j}(A, B) \quad \text{s.t. } BB' = \Sigma, B \in \mathcal{R}(\mu)$$

- ★ Nonlinear, possibly nondifferentiable transformations of  $\mu$ .
- ★ Obviously ...

$$\mathcal{I}_H^{\mathcal{R}}(\mu) \subseteq \times_{h=1}^H \left[ \underline{v}_{k_h, i_h, j_h}(\mu), \bar{v}_{k_h, i_h, j_h}(\mu) \right].$$

- ★ No need to assume the i.d. set is connected.

## PROJECTION REGION FOR $\lambda^H$

- ★ Let  $CS_T(1 - \alpha; \mu)$  be the (typical) Wald ellipsoid for  $\mu$ .
- ★ Let  $CS_T(1 - \alpha; \text{IRF}_{k,i,j})$  be the interval defined by:

$$\left[ \inf_{\mu \in CS_T(1-\alpha; \mu)} \underline{v}_{k,i,j}(\mu), \sup_{\mu \in CS_T(1-\alpha; \mu)} \bar{v}_{k,i,j}(\mu) \right]$$

- ★ The projection region for  $\lambda^H = \{\text{IRF}_{k_h, i_h, j_h}(A, B)\}_{h=1}^H$  is:

$$CS_T(1 - \alpha; \lambda^H) \equiv$$

$$CS_T(1 - \alpha; \text{IRF}_{k_1, i_1, j_1}) \times \dots \times CS_T(1 - \alpha; \text{IRF}_{k_H, i_H, j_H})$$

- ★ We now present the properties of  $CS_T(1 - \alpha; \lambda^H)$  as  $T \rightarrow \infty$

## **2. Assumptions and Results 1 to 4**

## RESULT 1: FREQUENTIST COVERAGE

- ★ Let  $P$  be a DGP for the data. Parameterized by  $(A, B, F)$ .
- ★ We want projection to be valid over a class  $\mathcal{P}$  of DGPs:
- ★ **A1:** Suppose the class of DGPs  $\mathcal{P}$  is such that

$$\liminf_{T \rightarrow \infty} \inf_{P \in \mathcal{P}} P\left(\mu(P) \in CS_T(1 - \alpha; \mu)\right) \geq 1 - \alpha.$$

- ★ **R1:** Under Assumption A1:

$$\liminf_{T \rightarrow \infty} \inf_{P \in \mathcal{P}} \inf_{\lambda^H \in \mathcal{I}_H^R(\mu(P))} P\left(\lambda^H \in CS_T(1 - \alpha; \lambda^H)\right) \geq 1 - \alpha.$$

## PROOF: STRAIGHTFORWARD PROJECTION ARGUMENT

Suppose that  $H = 1$ . For any  $\lambda \in \mathcal{I}_{k,i,j}^{\mathcal{R}}(\mu(P))$  :

$$\begin{aligned}
 & P\left(\lambda \in \left[ \inf_{\mu \in CS_T(1-\alpha)} \underline{v}_{k,i,j}(\mu), \sup_{\mu \in CS_T(1-\alpha)} \bar{v}_{k,i,j}(\mu) \right]\right) \\
 & \geq \\
 & P\left(\underline{v}_{k,i,j}(\mu(P)), \bar{v}_{k,i,j}(\mu(P)) \in \left[ \inf_{\mu \in CS_T(1-\alpha)} \underline{v}_{k,i,j}(\mu), \sup_{\mu \in CS_T(1-\alpha)} \bar{v}_{k,i,j}(\mu) \right]\right) \\
 & \quad \left(\text{as } \mathcal{I}_{k,i,j}^{\mathcal{R}}(\mu(P)) \subseteq [\underline{v}_{k,i,j}(\mu(P)), \bar{v}_{k,i,j}(\mu(P))]\right) \\
 & \geq \\
 & P\left(\mu(P) \in CS_T(1-\alpha)\right).
 \end{aligned}$$

## ROBUST BAYES FRAMEWORK

- ★ Let  $P^*$  be a prior for the structural parameters  $(A, B)$ .  
( $F$  is now a fixed known distribution; we use  $\mathcal{N}(0, \mathbb{I}_n)$ )
- ★ Represent the prior  $P^*$  in terms of  $(P_\mu^*, P_{Q|\mu}^*)$ ,  $Q \equiv \Sigma^{-1/2}B$ .  
(Orthogonal reduced-form parameterization Arias et. al [2014])
- ★ Let  $\mathcal{P}(P_\mu^*)$  denote the class of priors such that  $\mu \sim P_\mu^*$ .
- ★ The *robust credibility* of  $CS_T(1 - \alpha, \lambda^H)$  is defined as:

$$\inf_{P^* \in \mathcal{P}(P_\mu^*)} P^* \left( \lambda^H(A, B) \in CS_T(1 - \alpha; \lambda^H) \mid \mathbf{Y}_T \right)$$



## RESULT 2: ROBUST BAYESIAN CREDIBILITY

- ★ We can view robust credibility as a random variable (as it depends on the data  $\mathbf{Y}_T$ )
- ★ **A2:** Suppose that  $P^*$  is such that whenever  $\mathbf{Y}_T \sim f(\mathbf{Y}_T|\mu_0)$ :

$$P^*(\mu(A, B) \in CS_T(1 - \alpha; \mu) \mid \mathbf{Y}_T) = 1 - \alpha + o_p(\mathbf{Y}_T|\mu_0).$$

- ★ This is implied by the Bernstein von-Mises Theorem for  $\mu$ .
- ★ **R2:** Under Assumption 2:

$$\inf_{P^* \in \mathcal{P}(\mathcal{P}_\mu^*)} P^* \left( \lambda^H \in CS_T(1 - \alpha; \lambda^H) \mid \mathbf{Y}_T \right) \geq 1 - \alpha + o_p(\mathbf{Y}_T|\mu_0)$$

- ★ PROOF: Another embarrassingly simple projection argument!

## CALIBRATED PROJECTION

- ★ Yes: We know that projection inference is conservative!  
(both in terms of frequentist coverage and a robust credibility)
- ★ In theory, it is conceptually simple to remove ‘projection bias’  
(project a smaller Wald ellipsoid as suggested by KMS[2016])
- ★ In practice, removing the excess of robust Bayesian credibility is much easier than removing the excess of frequentist coverage.
- ★ We suggest an algorithm to ‘calibrate’ robust credibility.
- ★ The algorithm also removes the excess of frequentist coverage  
(provided the bounds of i.d. set are differentiable)

### RESULT 3: CALIBRATED ROBUST CREDIBILITY

- ★ Our calibration algorithm is based on the following result.
- ★ Suppose there is a nominal level  $1 - \alpha^*(\mathbf{Y}_T)$  s.t:

$$P_{\mu}^* \left( \times_{h=1}^H [\underline{v}_{k_h, i_h, j_h}(\mu), \bar{v}_{k_h, i_h, j_h}(\mu)] \subseteq CS_T(1 - \alpha^*(\mathbf{Y}_T), \lambda^H) \mid \mathbf{Y}_T \right) = 1 - \alpha$$

- ★ **R3:** Then, for every data realization:

$$\inf_{P^* \in \mathcal{P}(P_{\mu}^*)} P^* \left( \lambda^H(A, B) \in CS_T(1 - \alpha^*(\mathbf{Y}_T); \lambda^H) \mid \mathbf{Y}_T \right) = 1 - \alpha,$$

- ★ **PROOF:** slightly more involved. See Appendix A.2.

## CALIBRATION ALGORITHM

- ★ Take  $M$  draws ( $\mu_m^*$ ) from the posterior distribution of  $\mu$  (or from its asymptotic approximation based on BvM)
- ★ For each  $h = 1, \dots, H$  and each  $m = 1, \dots, M$  evaluate:

$$[\underline{v}_{k_h, i_h, j_h}(\mu_m^*), \bar{v}_{k_h, i_h, j_h}(\mu_m^*)]$$

(we use nonlinear numerical solvers to evaluate the bounds)

- ★ Fix a confidence level  $1 - \alpha_S < 1 - \alpha$ . Count how often:

$$[\underline{v}_{k_h, i_h, j_h}(\mu_m^*), \bar{v}_{k_h, i_h, j_h}(\mu_m^*)] \subseteq \text{CS}_T(1 - \alpha_S; \lambda_{k_h, i_h, j_h})$$

for all  $h = 1, \dots, H$ .

- ★ If RBC is not in between  $[1 - \alpha - \eta, 1 - \alpha + \eta]$ , change  $\alpha_S$ . ( $\eta$  is a tolerance level for the excess of robust credibility)

## RESULT 4: COVERAGE OF CALIBRATED PROJECTION

★ Question: Suppose we have found  $\alpha^*(Y_T)$ . Is it true that:

$$P_{\mu_0}([ \underline{v}_h(\mu_0), \bar{v}_h(\mu_0) ]) \subseteq \text{CS}_T(1 - \alpha^*(Y_T); \lambda_h), \forall h = 1, \dots, H \rightarrow 1 - \alpha ?$$

★ Answer: Yes. The following regularity conditions suffice:

★  $\underline{v}_h(\mu_0), \bar{v}_h(\mu_0)$  are differentiable at  $\mu_0$  for each  $h = 1, \dots, H$ .

★  $\sqrt{T}(\hat{\mu} - \mu_0) \xrightarrow{d} N(0, \Omega)$ ,  $\hat{\Omega}_T \xrightarrow{p} \Omega$ , and BvM holds.

★ PROOF: Under differentiability  $\text{CS}_T(1 - \alpha^*(Y_T); \lambda_h)$  is approximately:

$$\left[ \underline{v}_h(\hat{\mu}_T) - \frac{r_T^* \underline{\sigma}_h(\mu_0)}{\sqrt{T}}, \bar{v}_h(\hat{\mu}_T) + \frac{r_T^* \bar{\sigma}_h(\mu_0)}{\sqrt{T}} \right]$$

### **3. Implementation and illustrative example**

## PROJECTION AS AN OPTIMIZATION PROBLEM

- ★ To implement projection we need to solve the program:

$$\sup_{\mu \in \text{CS}_T(1-\alpha; \mu)} \bar{v}_{k,i,j}(\mu)$$

- ★ The confidence set for the reduced-form parameters is taken as:

$$\text{CS}_T(1-\alpha) \equiv \left\{ \mu \in \mathbb{R}^d \mid T(\hat{\mu}_T - \mu)' \hat{\Omega}_T^{-1} (\hat{\mu}_T - \mu) \leq \chi_{1-\alpha, d}^2 \right\}$$

- ★ Thus, the program of interest becomes:

$$\sup_{\mu \in \text{CS}_T(1-\alpha)} \left( \sup_{B \in \mathcal{I}_{k,i,j}^{\mathcal{R}}(\mu)} e_i' C_k(A) B_j \right)$$

- ★ Which is a twice differentiable, non-convex, nonlinear program:

$$\sup_{\mu, B} e_i' C_k(A) B_j \quad \text{s.t. } B \in \mathcal{I}_{k,i,j}^{\mathcal{R}}(\mu) \text{ and } \mu \in \text{CS}_T(1-\alpha)$$

## SOLUTION ALGORITHM

- ★ State-of-the-art nonlinear, non-convex, large-scale problems?
- ★ Large literature on local and global optimization algorithms  
(Local: AL, SQP, IP; Global: Multistart, GSearch, GAs)
- ★ Since the problem is non-convex, local solvers are not enough.  
(we use a two-phase solution algorithm: local+global)
- ★ SQP/IP 'most powerful local algorithm for nonlinear prog.'  
(Nocedal and Wright [2006], p. 253; implemented in `fmincon`)
- ★ For the global stage we use `Multistart`, `GlobalSearch`, `ga`  
(take the local solution as an input for global algorithm)



## EXAMPLE: DEMAND-SUPPLY SVAR BH(2015), ECMA

- ★ Effect of a structural shock on labor demand over wages/emp?  
(2-SVAR, 6 lags (1-AIC, 1-BIC), Q1-70/Q2-14:  $\Delta w_t, \Delta \eta_t$ .)

$$\begin{pmatrix} \Delta w_t \\ \Delta \eta_t \end{pmatrix} = A_1 \begin{pmatrix} \Delta w_{t-1} \\ \Delta \eta_{t-1} \end{pmatrix} + \dots + A_6 \begin{pmatrix} \Delta w_{t-6} \\ \Delta \eta_{t-6} \end{pmatrix} + B \begin{pmatrix} \epsilon_t^d \\ \epsilon_t^s \end{pmatrix}.$$

- ★ Sign restrictions set-identify structural demand and supply shocks:

$$B \equiv \begin{pmatrix} b_1 & b_3 \\ b_2 & b_4 \end{pmatrix} \quad \text{satisfies} \quad \begin{bmatrix} + & - \\ + & + \end{bmatrix}.$$

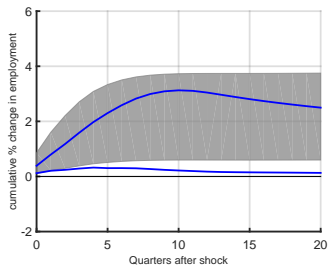
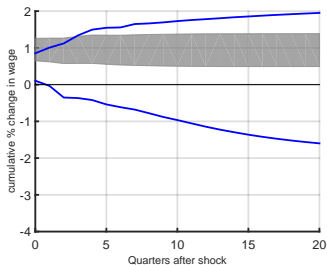
- ★ Elasticities of supply and demand:  $\alpha \equiv b_2/b_1$      $\beta \equiv b_4/b_3$ .

**Table:** ADDITIONAL IDENTIFYING RESTRICTIONS

Motivation	BH(2015)	This paper
Empirical studies report $\alpha \in [.27, 2]$	$\alpha \sim \max\{.6 + .6t_3, 0\}$	$.27 \leq \alpha \leq 2$
Empirical studies $\beta \in [-2.5, -.15]$	$\beta \sim \min\{-.6 + .6t_3, 0\}$	$-2.5 \leq \beta \leq -.15$
$\gamma = 0$ is too strong	$\gamma \sim \mathcal{N}(0, V)$	$-2V \leq \gamma \leq 2V$

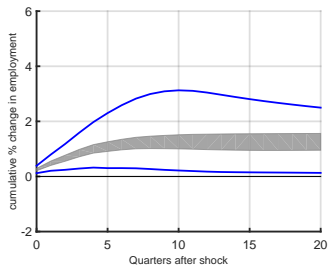
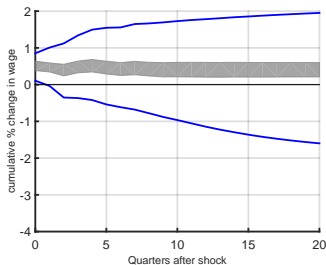
$$\gamma \equiv e_2'(\mathbb{I}_n - A_1 - A_2 - \dots - A_6)^{-1}B_1$$

68% PROJECTION REGION AND 68% CREDIBLE SET.



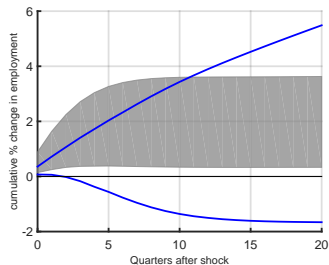
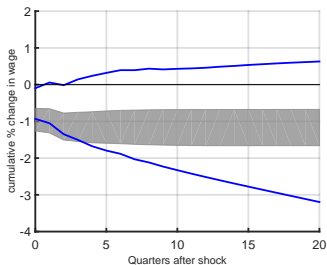
Expansionary Demand Shock  
(BH priors)

68% PROJECTION REGION AND 68% CREDIBLE SET.



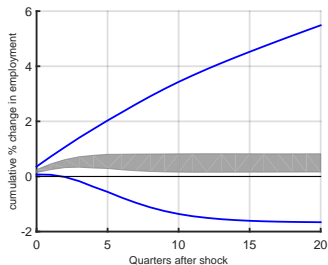
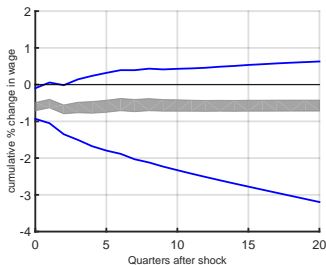
Expansionary Demand Shock  
(Uhlig's priors)

68% PROJECTION REGION AND 68% CREDIBLE SET.



Expansionary Supply Shock  
(BH priors)

68% PROJECTION REGION AND 68% CREDIBLE SET.

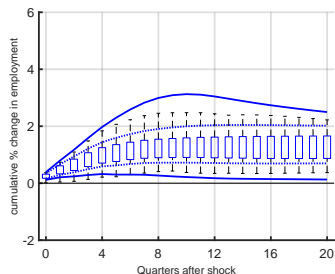
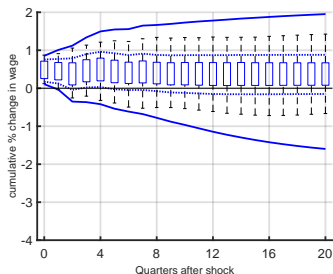


Expansionary Supply Shock  
(Uhlig's priors)

## ROBUSTNESS OR CONSERVATIVENESS?

- ★ Projection is informative about the effects of  $\epsilon^d$  on  $\eta_t \dots$
- ★ But not very informative about the rest of the dynamic effects.
- ★ This could be a consequence of the 'robustness' of projection,
- ★ Or a consequence of its conservativeness ( $> 1 - \alpha$ ).
- ★ To separate these effects, we report the calibrated projection.  
(represented as dotted line in the following figures)

## 68% PROJECTION REGION AND 68% CALIBRATED PROJECTION

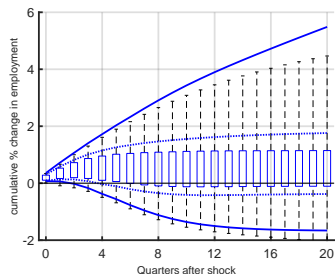
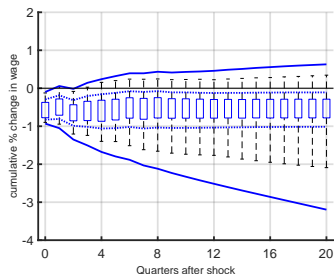


### Expansionary Demand Shock

Boxes: horizon-by-horizon 68% robust Bayesian credible set  
 Whiskers: minimum/maximum IRF (100,000 draws)



## 68% PROJECTION REGION AND 68% CALIBRATED PROJECTION



### Expansionary Supply Shock

Boxes: horizon-by-horizon 68% robust Bayesian credible set  
 Whiskers: minimum/maximum IRF (100,000 draws)

## COMMENTS

- ★ Credible sets differ substantially depending on the prior beliefs.  
(compare BH with Uhlig priors)
- ★ Prior-free, 'projected' region: qualitatively different inference.  
(only the employment response to demand shock significant)
- ★ SQP/IP: 12 min; Uhlig: 38 min; BH: 66min; Global: 9hrs.  
(see Table III, p. 20 in the paper)
- ★ Global algorithms do not improve the local solution.  
(see Appendix B, p. 48 in the paper)
- ★ Calibration of RCS takes around 3 minutes for  $M = 1,000$ .  
(and around 5 hours with  $M = 100,000$  and 50 parallel workers)

## 4. Conclusion

## MAIN MESSAGES FROM OUR PAPER

- ★ We studied the properties of projection inference for SVARs.  
(delivers frequentist and Robust Bayes interpretation)
- ★ We have emphasized the generality of projection inference.  
(can handle typical applications in applied macro work)
- ★ We thought seriously about computational feasibility.  
(implementation requires solving two mathematical programs)
- ★ We showed how to calibrate the RB credibility of projection.  
(which also calibrates coverage under some reg. conditions)

## BOTTOM LINE

We think that projection is a simple way to conduct inference in set-id SVARs.

(it has both frequentist coverage and Robust Bayes credibility)

(it is also general, feasible, and delivers simultaneous inference)

**Thanks very much for listening!**